

ERASMUS, BSc course

Winter 2020/2021

FPM&H, December 10, part II

Subject . Central Limit Theorem and its applications.

Introduction .

From the 1st part we know that if we have the sequence $(X_n)_{n \geq 1}$, where

$n \geq 1$ $d(X_n) = d(X)$ for some X with second moment

and (X_n) are i.i.d., then

$$\frac{1}{n} (X_1 + X_2 + \dots + X_n) \xrightarrow{\text{g.s.}} EX = m$$

What can happen if we consider the stronger type of convergence, namely with respect of distribution?

The Central Limit Theorem (CLT) gives the answer to that question.

Assumptions

Let X and (X_n) as above and
in addition, $\sigma_X^2 = \text{var}(X) > 0$ (so $X \neq \text{const}$).

Let us take:

$$1) \quad Y_n \stackrel{\text{def}}{=} \frac{1}{n} (X_1 + X_2 + \dots + X_n) \quad (\text{the average})$$

Then

$$E Y_n = \frac{1}{n} \cdot n E X = m,$$

$$\text{var}(Y_n) = \frac{1}{n^2} \cdot n \text{var}(X) = \frac{\sigma^2}{n} \quad (\sigma^2 = \sigma_X^2).$$

2) After standardization procedure of Y_n
we get for $n \geq 2$ and by the above

$$Z_n \stackrel{\text{def}}{=} \frac{Y_n - E Y_n}{\sqrt{\text{var}(Y_n)}} = \frac{Y_n - m}{\frac{\sigma}{\sqrt{n}}}$$

then C.L.T. says that:

$$Z_n \xrightarrow{d} \mathcal{N}, \text{ where } \mathcal{N} \in \mathcal{N}(0,1)$$

so we have

$\forall t \in \mathbb{R}$

$$F_{Z_n}(t) = P(R_{n,t} \cap \{Z_n \leq t\}) \longrightarrow$$

$$\Phi(t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-u^2/2} du$$

REMARKS

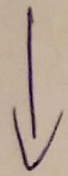
1) For sufficient large $n > 2$ we have
an approximation rule

$$F_{Z_n} \approx \Phi$$

2) We will check what this means for X :

We know from the above that

$$F_{Z_n}(h) = P\left(\sum_{i=1}^n U_i \leq m + t\right) =$$



$$\Phi(t) = P\left(\sum_{i=1}^n U_i \leq m + t \frac{\sigma}{\sqrt{n}}\right) =$$

$$= F_{Y_n}\left(m + t \frac{\sigma}{\sqrt{n}}\right), \text{ for } t \in \mathbb{R}.$$

So, by C.L.T. we can write

$$F_{Y_n}\left(m + t \frac{\sigma}{\sqrt{n}}\right) \approx \Phi(t) \quad (*)$$

Let $u = m + t \frac{\sigma}{\sqrt{n}} \Rightarrow t = \frac{u-m}{\frac{\sigma}{\sqrt{n}}}$, and (*) means that

$$F_{Y_n}(u) = \Phi\left(\frac{u-m}{\frac{\sigma}{\sqrt{n}}}\right).$$

But

$$\Phi\left(\frac{u-m}{\frac{\sigma}{\sqrt{n}}}\right) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\frac{u-m}{\frac{\sigma}{\sqrt{n}}}} e^{-t^2/2} dt.$$

let us apply the substitution rule in
the integral:

$$t = \frac{y-m}{\frac{\sigma}{\sqrt{n}}} \quad dt = \frac{\sqrt{n}}{\sigma} dy, \text{ so}$$

$$\Phi\left(\frac{y-m}{\frac{\sigma}{\sqrt{n}}}\right) = \frac{1}{\sqrt{2\pi}} \frac{\sqrt{n}}{\sigma} \int_{-\infty}^t e^{-\left(\frac{y-m}{\frac{\sigma}{\sqrt{n}}}\right)^2 / 2} dy$$

$$= \underbrace{\frac{1}{\sqrt{2\pi} \frac{\sigma}{\sqrt{n}}}}_{\text{the cumulative dist. function corresponding to } N(m, \frac{\sigma^2}{n})} \int_{-\infty}^t e^{-\frac{(y-m)^2}{2 \frac{\sigma^2}{n}}} dy$$

the cumulative dist. function
corresponding to $N(m, \frac{\sigma^2}{n})$.

So, finally $\Phi_{\sqrt{in}} \approx N(m, \frac{\sigma^2}{n})$

Let us consider the following situation:

$$\text{for } n \geq 2 \quad a < b, \quad B_n \in \mathcal{B}(n, p) \\ p \in (0, 1)$$

$$A_{a,b,n} = \{ \omega : a \leq B_n(\omega) < b \}$$

and let us try to calculate $P(A_{a,b,n})$.

We will get ~~one by one~~: successively after averaging

$$P(A_{a,b,n}) = P\left(\left\{ \omega : \frac{a}{n} \leq \frac{B_n(\omega)}{n} < \frac{b}{n} \right\}\right)$$

$$\text{But } E\left(\frac{B_n}{n}\right) = p, \quad \text{Var}\left(\frac{B_n}{n}\right) = \frac{pq}{n}$$

so after standardization procedure we get

$$= P\left(\left\{ \frac{\frac{a}{n} - p}{\sqrt{\frac{pq}{n}}} \leq \frac{\frac{B_n(\omega)}{n} - p}{\sqrt{\frac{pq}{n}}} < \frac{\frac{b}{n} - p}{\sqrt{\frac{pq}{n}}} \right\}\right)$$

Finally, by C.L.T. we can write

$$P(A_{a,b,n}) \approx \Phi\left(\frac{b}{\sqrt{pq}} - p\right) - \Phi\left(\frac{a}{\sqrt{pq}} - p\right).$$

Remark.

The last formula is a first version of
 C.L.T - it is called de Moivre-Laplace
Theorem (from XVIII c.)

Example.

Let $n=100$, $a=30$, $b=70$, $p=0,25$

Then

$$P(A_{a,b,n}) \approx \Phi(10,392) - \Phi(1,1537)$$

$$\approx 1 - \Phi(1,1537) =$$

$$= 1 - 0,8943 = \underline{\underline{0,1057}}$$

Example.

We will estimate

$$(*) \quad P(\text{given: } B_n(u) = k) \quad \text{for } n=100, \\ k=20 \\ p=0,25$$

We note, that

$$(*) = P(\text{given: } 29,5 \leq B_n(u) < 30,5)$$

$$\approx \Phi(1,2707) - \Phi(1,0392)$$

$$= \underline{\underline{0,0495}}$$

On the other hand we have:

$$P(\text{given: } B_{100}(u) = 20) = \binom{100}{20} (0,25)^{20} (0,75)^{80}$$

But:

$$\binom{100}{30} = 2,9372 \cdot 10^{25}$$

$$(0,25)^{30} = 8,6736 \cdot 10^{-19}$$

$$(0,75)^{70} = 1,7959 \cdot 10^{-9}$$

and finally

$$P(\text{given: } P_{\text{no}}/N = 30) = \underline{0,045}$$

If we compare this result we get

$$|0,0495 - 0,045| = \underline{\underline{0,004}}$$

SO IT WORKS !!!