

ERASMUS B3 course

Winter 2020/2021

FPM&IT November 19

Subject 1^o The concept of P.D. - continuous case - ~~Examples~~
2^o The parameters of P.D.

We begin with the following example

Ex 1. Let $X \in U([0, 1])$ and $\lambda > 0$.

We define $Y = -\frac{1}{\lambda} \ln(1 - X)$. We will find the distribution of Y .

By assumption there exists at least one RE and
is KPM (Ω, Σ, P) , such that

$$F_X(t) = P(\exists \omega \in \Omega: X(\omega) < t) \text{ and}$$

$$Y(\omega) = -\frac{1}{\lambda} \ln(1 - X(\omega)). \text{ Since } X(\omega) \in (0, 1)$$

the last definition is correct.

Moreover, $\forall \omega \in \Omega$ $1 - X(\omega) \in (0, 1)$, so

$\ln(1 - X(\omega)) < 0$, and consequently $Y(\omega) > 0$.

It means, that for $t \leq 0$

$$F_Y(t) = 0.$$

①

Therefore, to find F_Y it suffices to assume $t > 0$.

For $t > 0$ we have

$$F_Y(t) = P(\text{Survivor: } Y(t) < t) = P(\text{Survivor: } -\frac{1}{\lambda} \ln(1-X(t)) < t)$$

$$= P(\text{Survivor: } \ln(1-X(t)) > -\lambda t) =$$

$$= P(\text{Survivor: } 1-X(t) > e^{-\lambda t}) =$$

$$= P(\text{Survivor: } X(t) < 1-e^{-\lambda t}) = F_X(1-e^{-\lambda t}).$$

So,

$$F_Y(t) = \begin{cases} 0, & t \leq 0 \\ F_X(1-e^{-\lambda t}), & t > 0, \text{ and} \end{cases}$$

$$f_Y(t) = \begin{cases} 0, & t \leq 0 \\ \frac{d}{dt} F_X(1-e^{-\lambda t}), & t > 0 \end{cases} =$$

$$= \begin{cases} 0, & t \leq 0 \\ f_X(1-e^{-\lambda t}) e^{-\lambda t} \cdot \lambda, & t > 0 \end{cases}$$

(2)

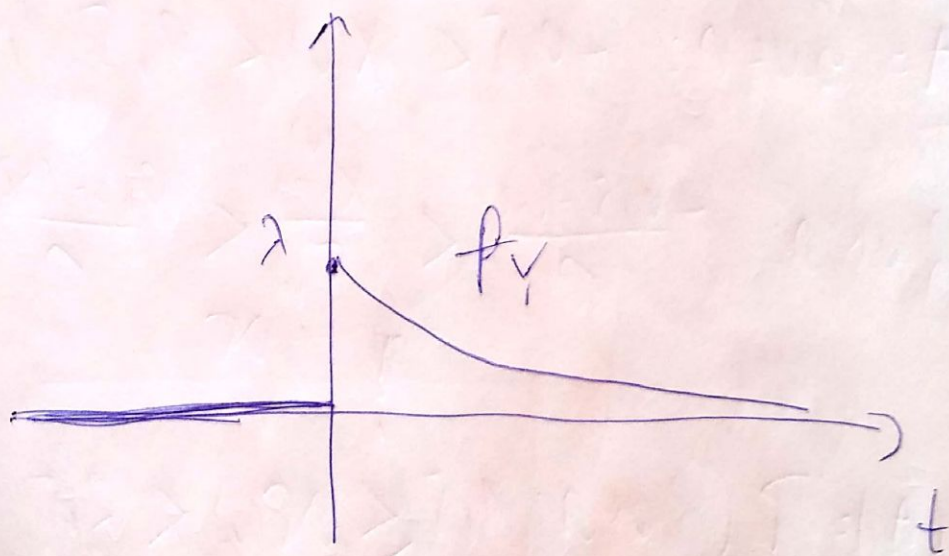
Then

$$f_X(u) = \begin{cases} 1, & u \in [0, 1] \\ 0, & u \notin [0, 1] \end{cases}$$

and $u = 1 - e^{-\lambda t} \in [0, 1] \Leftrightarrow t > 0,$

so finally

$$f_Y(t) = \begin{cases} 0, & t \leq 0 \\ \lambda e^{-\lambda t}, & t > 0 \end{cases}$$



Def 1 We write that $X \in W(\lambda), \lambda > 0$, if

$$f_X(t) = \begin{cases} 0, & t \leq 0 \\ \lambda e^{-\lambda t}, & t > 0 \end{cases}$$

X is called an exponential distribution with param.

(3)

λ

No we find the F_X for $X \sim U(x)$.

From general theory

$$F_X(t) = \int_{-\infty}^t f_X(u) du = \begin{cases} 0 & t \leq 0 \\ t f_X(t) + \int_{-\infty}^t f_X(u) du, & t > 0 \end{cases}$$

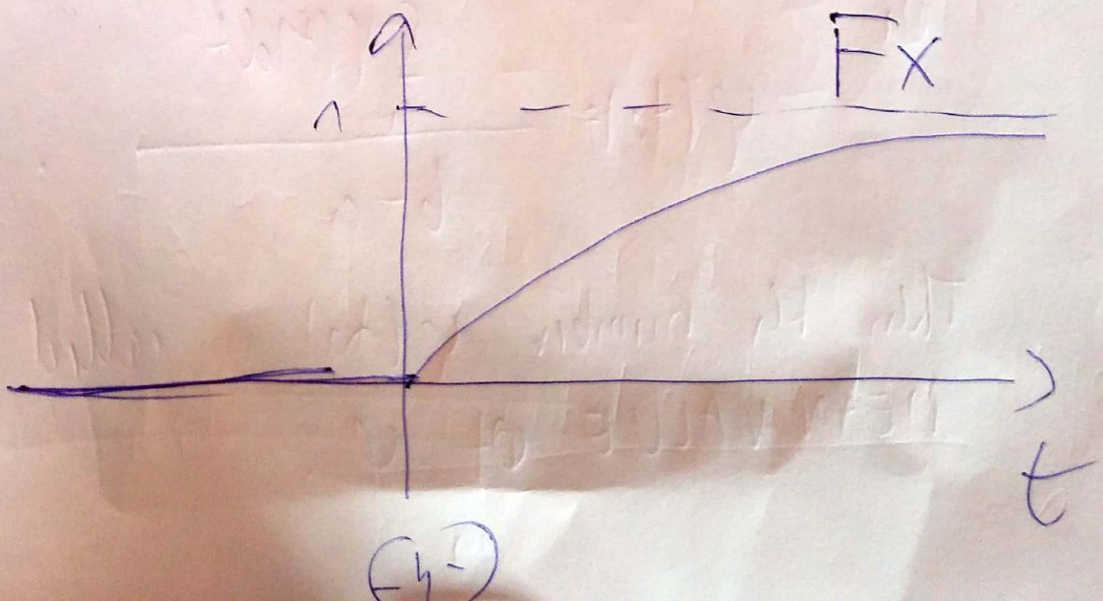
BWA

$$\int_{-\infty}^t x e^{-\lambda x} dx = \int_{-\infty}^0 0 + \int_0^t x e^{-\lambda x} dx =$$

$$= x \left[-\frac{1}{\lambda} e^{-\lambda x} \right]_0^t = - \left[e^{-\lambda t} - 1 \right] = 1 - e^{-\lambda t}$$

and

$$F_X(t) = \begin{cases} 0, & t \leq 0 \\ 1 - e^{-\lambda t}, & t > 0 \end{cases}$$



Gaussian distribution

We say that $X \in N(m, \sigma^2)$, where $m \in \mathbb{R}$, $\sigma > 0$
if the density function f_X is given as follows

$$f_X(t) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(t-m)^2}{2\sigma^2}}, \quad t \in \mathbb{R},$$

where "e" denotes the Euler's number.

STATEMENT

Let $X \in N(m, \sigma^2)$. Then

$$Y = \frac{X-m}{\sigma} \in N(0, 1) \quad (\text{so } m=0, \sigma=1).$$

Proof

$$\begin{aligned} F_Y(t) &= P(\text{event: } Y(t) < t) = P(\text{event: } \frac{X(t)-m}{\sigma} < t) \\ &= P(\text{event: } X(t) < m + t\sigma) = F_X(m + t\sigma). \end{aligned}$$

Then

$$f_Y(t) = \frac{d}{dt} F_Y(t) = F_X'(m + t\sigma) \cdot (m + t\sigma)' =$$

$$= f_X(m+t\sigma) \cdot \sigma$$

Therefore by def. of f_X we have

$$f_Y(t) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(m+t\sigma-m)^2}{2\sigma^2}} \cdot \sigma =$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} = f_X(t) \Big|_{\substack{m=0 \\ \sigma=1}}$$

and $Y \in N(0,1)$.

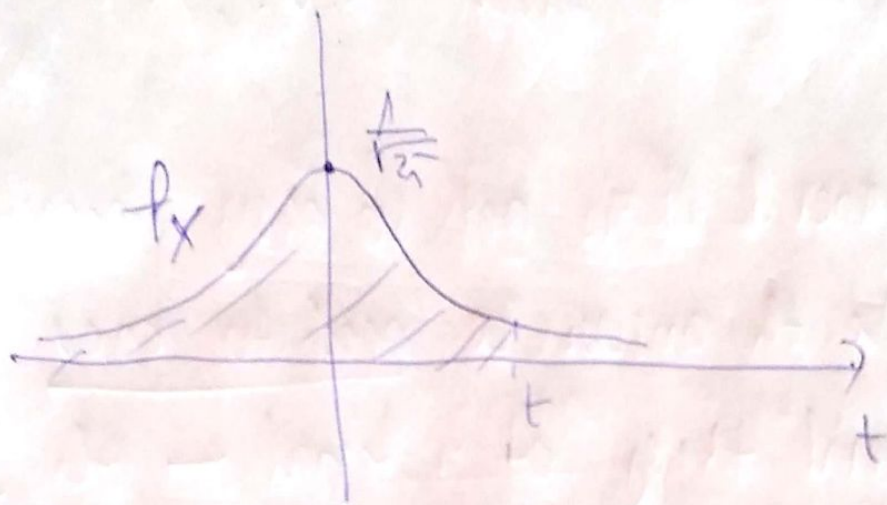
If $X \in N(\mu, \sigma^2)$ we say that we have a standard gaussian distribution. Then the

C.P.D.F. we denote by Φ .

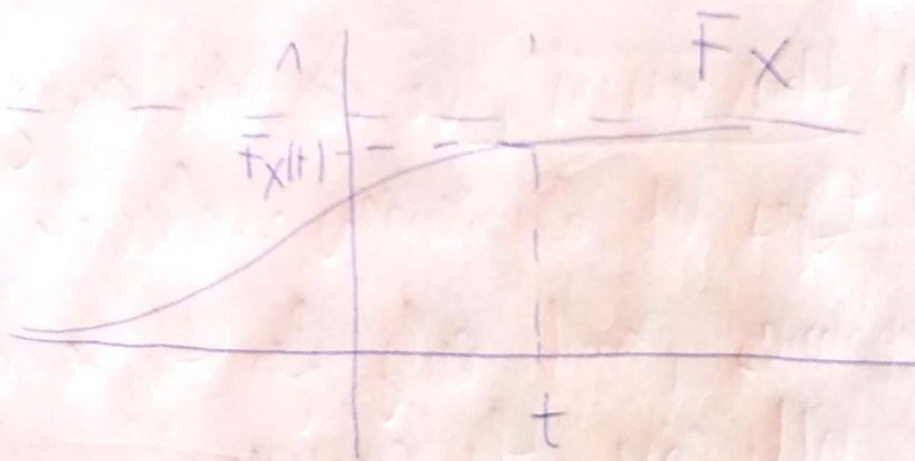
Remarks concerning $N(0,1)$

- 1). The procedure realized in the STATEMENT is called "standardization procedure"

2) the graph of $f_X(t) = \frac{1}{\sqrt{2\pi}} e^{-t^2/2}$, $t \in \mathbb{R}$



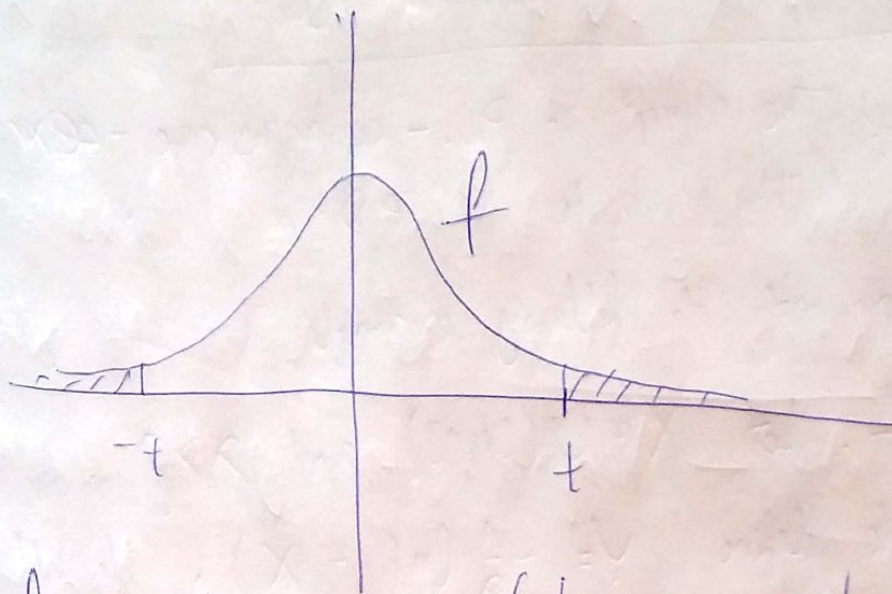
is called "the bell curve"



Then $F_X(t) = P(\text{Rv} \leq X(t)) = \int_{-\infty}^t f_X(u) du$

In particular, $\Phi(0) = 0.5$

⑤ Fix $t > 0$ as above



Then for $-t$ we have (by symmetry)

$$\bar{\Phi}(-t) = P(\text{Re} \in \mathcal{A}: X(t) < -t)$$

||

+

$+\infty$

$$\int_{-\infty}^t f_X(u) du = \int_t^{+\infty} f_X(u) du = P(\text{Re} \in \mathcal{A}: X(t) \geq t)$$

and finally

$$\bar{\Phi}(-t) = 1 - \bar{\Phi}(t)$$

$t > 0$

so $\bar{\Phi}$ is determined by positive arguments.

TASK 1

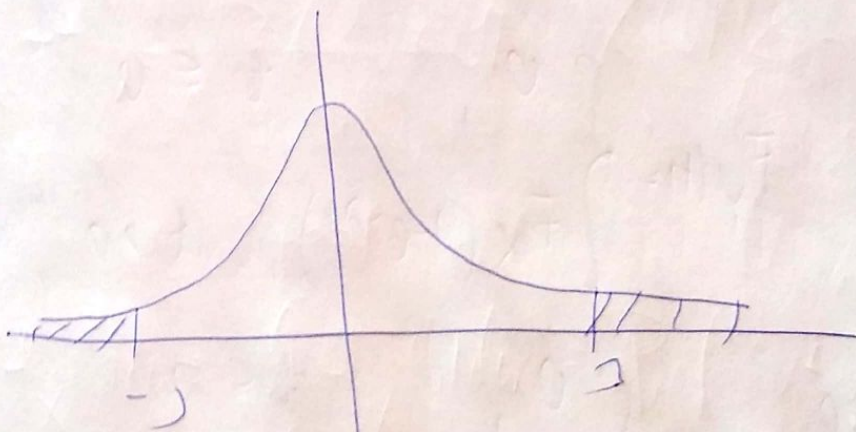
Compute

$$P(\text{Zwei: } |X(t)| < t), \quad t > 0$$

$$X \in N(0, 1)$$

b) Later we will show, that für $t_0 = 3$

$$P(\text{Zwei: } X(t) \geq 3) < 0,01$$



so almost all information concerning of $X \in N(0, 1)$ is included in $[-3, 3]$.

5) In every (well written) book you can find the table of $N(0, 1)$.

Ex 2

For $X \in N(-2, 1)$, find $P(A)$,

where

$$A = \{ \omega \in \Omega : |X(\omega)| < 0,5 \}$$

We use the standardization procedure

$$X \longrightarrow \frac{X - \mu}{\sigma} \Big|_{\substack{\mu = -2 \\ \sigma = 1}} \in N(0, 1)$$

$$A = \{ \omega \in \Omega : -0,5 < X(\omega) < 0,5 \}$$

$$= \{ \omega \in \Omega : -\frac{0,5 - (-2)}{1} < \frac{X(\omega)}{1} < \frac{0,5 - (-2)}{1} \}$$

so

$$P(A) = P(\{ \omega \in \Omega : 1,5 \leq Z(\omega) < 2,5 \})$$

$$= \Phi(2,5) - \Phi(1,5)$$

from the TABLE: 0,99379

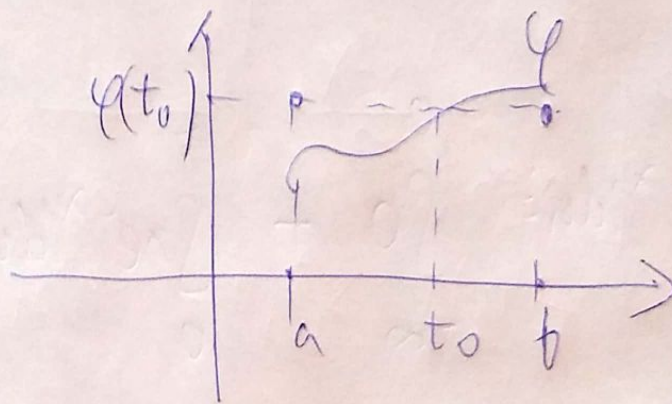
0,93319

An introduction to the concept of the parameters

I The concept of the MEAN VALUE

We begin from the same remark.

Suppose that we have a continuous and positive function φ on finite interval $[a, b]$



From the theory of integral it follows that

$$\exists t_0 \in (a, b) \text{ that } \int_a^b \varphi(t) dt = (b-a) \varphi(t_0),$$

$$\text{or } \varphi(t_0) = \frac{\int_a^b \varphi(t) dt}{b-a}.$$

Then the number $\varphi(t_0)$ is called the MEAN VALUE of φ .

Note, let the measure of the set

$$\{(x,y) \in \mathbb{R}^2 : x \in [a,b], 0 \leq y \leq \varphi(x)\}$$

is the same as a rectangle ~~with~~ with the side measure: $b-a$ and $\varphi(b)$

Now, we are going to adopt the above idea to the case of probability distribution.

So let X be a given R.V., and (Ω, Σ, P) a corresponding K.P.M., so

$$X: \Omega \longrightarrow \mathbb{R}$$

~~The~~ Analogous to the ~~old~~ classical ~~problem~~ MEAN VALUE PROBLEM we define

Def 2 If there exists the integral

$$\int_{\Omega} X dP, \text{ then we say that } X \text{ has}$$

its mean value m or has an expected
value EX and we write

$$m = EX = \int X dP.$$

Th 1 (Linear property) -

The operation (as an integral)

$X \longrightarrow EX$ has linear

property, so:

If X_1, X_2 have mean value, then

for all $a, b \in \mathbb{R}$,

$aX_1 + bX_2$ has mean value

and we have

$$E(aX_1 + bX_2) = aEX_1 + bEX_2$$

and $Ec = c$ ($c \in \mathbb{R}$ constant).

Now it is a time to explain how to determine
 $\int X dP$ for given R.V. X .

In general it is not so easy, because
to give the answer we need to know the
general theory of integral — the Lebesgue's
integral.

For this reasons, we consider only of
two types of distributions — discrete and
continuous, and only for those cases
we will give the answer.